Convex polytopes, interacting particles, spin glasses, and finance

Sourav Chatterjee (UCB)

joint work with

Soumik Pal (Cornell)

- Systems of particles governed by joint stochastic differential equations.
- Drift of each particle depends on its relative position with respect to other particles.
- Continuous drifts: classical case (McKean-Vlasov).
- Completely different problem if drifts are prone to abrupt changes.

- X_1, \ldots, X_n are *n* diffusive particles, starting from zero.
- At each point of time, the lowest one gets a fixed upward drift.

Formally,

$$dX_i(t) = \delta_i(t)dt + dB_i(t),$$

where

$$\delta_i(t) = egin{cases} 1 & ext{if } X_i(t) = \min_j X_j(t), \ 0 & ext{otherwise}, \end{cases}$$

and B_1, \ldots, B_n are independent Brownian motions.

 Occurs in finance (Banner, Fernholz & Karatzas '05). Uses reflecting Brownian motions and the Harrison-Williams machinery. More later.

- Again, X_1, \ldots, X_n are *n* Brownian particles starting from zero.
- At any point of time, the particle that is farthest from zero has a unit drift towards zero.
- Same definition in \mathbb{R}^d .
- Question: What is the stationary distribution of this system?
- Cannot use RBM tools in $d \ge 2$.

A snapshot from the stationary law (n = 10,000)



- Generate U_1, \ldots, U_n i.i.d. uniformly from unit ball in \mathbb{R}^d .
- Let $V_i = U_i / \max_j \|U_j\|$, i = 1, ..., n.
- Generate Γ ~ Gamma(nd, 1) independently.

Let

$$Y_i=\frac{1}{2}\Gamma V_i,\ i=1,\ldots,n.$$

Then (Y_1, \ldots, Y_n) follows the stationary distribution of our system.

We have

$$dX_i(t) = \delta_i(t)dt + dB_i(t),$$

where

$$\delta_i(t) = \begin{cases} 1 & \text{if } |X_i(t)| = \max_j |X_j(t)|, \ X_i(t) < 0, \\ -1 & \text{if } |X_i(t)| = \max_j |X_j(t)|, \ X_i(t) > 0, \\ 0 & \text{otherwise.} \end{cases}$$

Alternatively, $d\mathbf{X}(t) = -\nabla k(\mathbf{X}(t))dt + d\mathbf{B}(t),$ where $k(\mathbf{x}) = \max_i |x_i|.$

Proposition

Consider the s.d.e.

$$d\mathbf{X}(t) = -\nabla k(\mathbf{X}(t))dt + d\mathbf{B}(t),$$

where k is any absolutely continuous function. Assume that $\exp(-2k(\mathbf{x}))$ is integrable and

$$\int_{\mathbb{R}^n} \|\nabla k(\mathbf{x})\|^2 e^{-2k(\mathbf{x})} d\mathbf{x} < \infty.$$

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Then the probability distribution given by the un-normalized density $\exp(-2k(\mathbf{x}))$ provides a reversible, invariant probability distribution μ for the process $\mathbf{X}(t)$. Under some further conditions (...), the law of $\mathbf{X}(t)$ converges to μ in TV.

▶ So in our problem, *k* is the Minkowski norm of the unit cube, and the stationary density is $\propto e^{-2 \max_i |x_i|}$.

Lemma

If k is the Minkowski norm of a convex set $C \subseteq \mathbb{R}^n$, then picking from $\exp(-2k(\mathbf{x}))$ is the same as: Picking $\mathbf{U} \sim Unif(C) \rightarrow dividing$ by $2k(\mathbf{U}) \rightarrow multiplying$ by independent Gamma(n, 1).

- ▶ Proof is simple, using "polar coordinates" induced by *C*.
- Picking from Unif(C) is can be handled theoretically if C can be easily triangulated, e.g. if C is a simplicial polytope.

Each (n-1)-dimensional face is a simplex.



- $C = \bigcup_{j=1}^{d} C_j$; simplices C_j . • Let $\mathbf{X}_j \sim Unif(C_j)$.
- $\mathbb{P}(\Pi = j) = Vol(C_j)/Vol(C).$

$$\begin{split} \mathbf{X} &= \sum_{j=1}^{d} \mathbf{X}_{j} \ \mathbb{I}(\Pi = j) \sim \textit{Unif}(C). \\ & \frac{\Gamma}{2k(\mathbf{X})} \mathbf{X} = \frac{\Gamma}{2k(\mathbf{X}_{j})} \mathbf{X}_{j}, \quad \text{if } \Pi = j. \\ & \frac{\Gamma}{2k(\mathbf{X}_{j})} \mathbf{X}_{j} - \text{affine transformation of IID Exponentials.} \end{split}$$

► Recall:

$$dX_i(t) = \delta_i(t)dt + dB_i(t),$$

where

$$\delta_i(t) = egin{cases} 1 & ext{if } X_i(t) = \min_j X_j(t), \ 0 & ext{otherwise}. \end{cases}$$

- Here $k(\mathbf{x}) = -\min_i x_i$.
- ► This k is not the Minkowski norm of any convex set. Nor is exp(-2k(x)) integrable.
- The process does not converge to a stationary law.

Atlas model contd...

▶ Although $k(\mathbf{x}) = -\min_i x_i$ is not a Minkowski norm, it *is* so when restricted to $\sum_i x_i = 0$. On this hyperplane, it is the norm induced by a regular simplex.



Figure: Level sets of $k(\mathbf{x})$ and their intersection with $\{\mathbf{x} : \sum x_i = 0\}$

So, we project the process on to {x : ∑x_i = 0} and apply previous result. This gives the stationary law of

$$(X_1(t)-\bar{X}(t),\ldots,X_n(t)-\bar{X}(t)).$$

- ► Explicit description: Generate i.i.d. exponential r.v.'s with mean *n*/2 and subtract off the mean.
- Obtained by Banner-Fernholz-Karatzas '05, and Pitman (tech. report). Both use Harrison-Williams RBM tools.
- General rank-dependent drifts: Stated as a major open problem in BFK '05. Solved by Pitman (tech. report). Easy solution by the following general result...

Theorem (C. & Pal '06) Start with $k : \mathbb{R}^n \to \mathbb{R}$.

$$d\mathbf{X}(t) = -\nabla k(\mathbf{X}(t))dt + d\mathbf{B}(t).$$

Suppose there exists a subspace H such that

- ► $k(x) = k_1(y) + k_2(z),$ $y = P_H(x),$ z = x y.
- $k_1 \ge 0$, cont., positively homogenous.

•
$$\{x \in H: k_1(x) = 0\} = \{0\}.$$

Then

- Y(t) = P_HX(t) ⇒ unique stationary distribution. Density ∝ exp(-2k₁(y)). Can be generated using earlier trick.
- Exponentially fast. Reversible when stationary.

Example:
$$-\min x_i = -\min(x_i - \overline{x}) - \overline{x}, \quad H = \{\mathbf{x} : \overline{x} = 0\}.$$

Some finance

- Equity market with *n* stocks.
- ▶ For the *i*th stock (company), define
 - capital, S_i = number of outstanding shares \times share price
 - market weight,

$$\mu_i = \frac{\text{capital of } i\text{th stock}}{\text{total capital}} = \frac{S_i}{\sum_i S_j}.$$

- Capital distribution curve = log-log plot of μ versus rank.
 - Ordered market wts: $\mu_{(1)} \leq \mu_{(2)} \leq \ldots \leq \mu_{(n)}$.
 - Plot $\log k$ versus $\log \mu_{(n-k+1)}$.
- Economic theory (e.g. Simon '55) predicts that capital distribution curve should be a straight line.
- ▶ In reality, the curve is concave and remarkably stable in time.



- Jovanovic ('82), Hopenhayn ('92), Axtell ('99), Hashemi ('00), Kou & Kou ('01).
- ► However, they all assume that the market converges rapidly to equilibrium, which is never true in reality!
- BFK's Atlas model tries to correct that, but recovers classical curve (straight line).

- Log-capital: $X_i = \log S_i$.
- Black-Scholes: Log-capital of a single company is modeled as a Brownian motion with drift.
- ► Natural extension for whole market: Log-capitals of all companies modeled as interacting particles on ℝ.
- Empirical observation: Larger stocks have slower upward mobility than smaller stocks.
- One way to model this: The interacting particles (log-capitals) pull each other by a "gravitational force".

- Let $X_i(t)$, i = 1, ..., n be the log-capitals at time t.
- 'Gravitational force' between *i* and *j* is proportional to $sign(X_i X_j)$.
- Resulting model:

$$dX_i(t) = -\frac{\alpha}{n}\sum_{j=1}^n \operatorname{sign}(X_i(t) - X_j(t))dt + dB_i(t), \ i = 1, \ldots, n.$$

Represents a toy model of flow of capital from larger to smaller stocks. The parameter α determines the strength of the flow.

Capital distribution: Reality vs. Gravity model

Real data (left) vs. simulations from gravity model with $\alpha = 1/2$ (right).



Here

$$k(\mathbf{x}) = \frac{\alpha}{n} \sum_{i,j=1}^{n} |x_i - x_j|.$$

- k(x) is not the Minkowski norm of any convex set in ℝⁿ. Also, exp(-2k(x)) is not integrable.
- ► However, k(x) is the Minkowski norm of a polytope when restricted to H = {x : ∑x_i = 0}.
- The polytope is regular and simplicial. Uniform generation is easy to describe.

Stationary distribution

▶ **X**(*t*) does not converge in law. However,

$$(X_1(t)-\bar{X}(t),\ldots,X_n(t)-\bar{X}(t))$$

does converge to an equilibrium measure.

- Suppose (Y_1, \ldots, Y_n) is drawn from this limiting distribution. Let $Y_{(1)} \leq \cdots \leq Y_{(n)}$ denote the Y_i 's arranged in increasing order.
- ▶ Let $\Delta_i = Y_{(i+1)} Y_{(i)}$. Then $\Delta_1, \ldots, \Delta_{n-1}$ are independent, and

$$\Delta_i \sim Exp\left(\frac{2\alpha i(n-i)}{n}\right)$$

► Each possible ordering corresponds to one face of the polytope. Simplicial polytope ⇒ exponentials. Recall: Market weight of *i*th company is

$$\mu_i(t) = rac{ ext{capital of } i ext{th stock}}{ ext{total capital}} = rac{e^{X_i(t)}}{\sum_i e^{X_j(t)}}.$$

Market diversity, as defined by Fernholz ('99):

$$\mu_{(n)}(t) := \max_{i} \mu_{i}(t) \in [0, 1].$$

• Under the gravity model, this has a limiting distribution as $t \to \infty$.

Theorem

Consider the gravity model with strength parameter α :

$$dX_i(t) = -\frac{\alpha}{n}\sum_{j=1}^n \operatorname{sign}(X_i(t) - X_j(t))dt + dB_i(t), \ i = 1, \dots, n.$$

Let $\mu_{(n)}$ denote the diversity in equilibrium. Then as $n \to \infty$,

- If $\alpha > 1/2$, then $\mu_{(n)} \sim n^{-(2\alpha-1)/2\alpha}$. Diversity exists.
- If $\alpha = 1/2$ then $\mu_{(n)} \sim (\log n)^{-1}$. Diversity exists to a lesser extent.
- If $\alpha < 1/2$, then $\mu_{(n)} \not\rightarrow 0$. No diversity.

• $\mu_{(n)}$ can be written as

$$\frac{1}{1+e^{-\xi_1}+e^{-(\xi_1+\xi_2)}+\cdots+e^{-(\xi_1+\cdots+\xi_{n-1})}},$$

where $\xi_i \sim Exp(2\alpha i(n-i)/n).$
When $i \ll n$,
 $\mathbb{E}(\xi_1+\cdots+\xi_i) \sim \frac{\log i}{2\alpha}.$

Vague intuition:

$$\mu_{(n)} \sim \frac{1}{1 + 2^{-1/2\alpha} + 3^{-1/2\alpha} + \dots n^{-1/2\alpha}} \sim \begin{cases} n^{-(2\alpha-1)/2\alpha} & \text{if } \alpha > 1/2, \\ (\log n)^{-1} & \text{if } \alpha = 1/2, \\ \text{const.} & \text{if } \alpha < 1/2. \end{cases}$$

Rigorous proof via martingales and Poincaré inequalities.

- Let Σ be the set of possible configurations of a physical system.
- Let {Z_σ, σ ∈ Σ} be a fixed collection of i.i.d. gaussian random variables.
- Derrida's Random Energy Model (REM) assigns a probability measure on Σ by putting mass ∝ exp(βZ_σ) at each σ ∈ Σ.
- Exhibits phase transition as β varies. Mathematical reason is the same as for the gravity model phase transition.
- More complex models of spin glasses assume that the Z_σ's have some correlation structure. Similar to more complex versions of the gravity model.

A generalized gravity model

• Gravity model on a graph
$$G = (V, E)$$
:

$$dX_i(t) = -\alpha \sum_{j:(i,j)\in E} \operatorname{sign}(X_i(t) - X_j(t))dt + dB_i(t).$$

- ► As before, (X₁(t) X
 (t),...,X_n(t) X
 (t)) converges to a stationary law. Polytope is simplicial but not regular.
- ▶ Let $(Y_1, ..., Y_n)$ be a draw from the stationary distribution. Let $Y_{(1)} \le \cdots \le Y_{(n)}$ be the ordered sample.
- Let Π_i = the index j such that $Y_{(i)} = Y_j$.
- ▶ For each $\pi \in S_n$ and $1 \le i \le n-1$, let

$$f_i(\pi) = \#\{(j,k) : j \le i < k, (\pi_j, \pi_k) \in E\}.$$

- Then $\mathbb{P}(\Pi = \pi) \propto (\prod_{i=1}^{n-1} f_i(\pi))^{-1}$.
- Given $\Pi = \pi$, the increments $Y_{(i+1)} Y_{(i)}$ are independent and

$$Y_{(i+1)} - Y_{(i)} \sim Exp(2\alpha f_i(\pi)).$$